

Local Bayesian mixture of different occupations for nuclear mass predictions

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Outline

- ① Background
- ② Bayesian framework
- ③ Nuclear models
- ④ Results
- ⑤ Conclusion

Global nuclear mass predictions

Many different theoretical frameworks

- Macroscopic-microscopic finite range droplet model
- Density functional theory
- Empirical microscopic shell models

Important for applications

- Nucleosynthesis, r -process

Previous work

Recent use of Bayesian inference

- Bayesian model averaging and uncertainty quantification¹
 - Find most important/“best” model
 - Uncertainty of each model
- Bayesian model mixing²
 - Localized model importance
 - Improve rms and uncertainty quantification

Could these methods be combined to inform local nuclear properties?

¹Saito, Y., Dillmann, I., Kruecken, R., Mumpower, M. R., and Surman, R. (2024). [Uncertainty quantification of mass models using ensemble bayesian model averaging](#). *Physical Review C*, 109(5):054301

²Kejzlar, V., Neufcourt, L., and Nazarewicz, W. (2023). [Local bayesian dirichlet mixing of imperfect models](#). *Scientific Reports*, 13(1):19600

Bayesian model averaging

- Nuclear space $\mathcal{X} \ni x_i$, data D , masses predicted by models, $\mathcal{M}_1, \dots, \mathcal{M}_p$
- “Standard” BMA:

$$p(\Delta|D) = \sum_{k=1}^p p(\Delta|\mathcal{M}_k, D)p(\mathcal{M}_k|D)$$

- Weights: $w_k = p(\mathcal{M}_k|D) \geq 0$, $\sum w_k = 1$
- Assume normal probability distributions:

$$p(\Delta(x_i)|\mathcal{M}_k, D) = \phi\left(\frac{\Delta(x_i) - m_k(x_i)}{\sigma_k}\right)$$

Local model mixing

- Introduce hyperparameters

$$\mathbf{w}(x_i) | \boldsymbol{\alpha}(x_i) \sim \text{Dirichlet}(\boldsymbol{\alpha}(x_i))$$

- Gaussian process gives local dependence:

$$\alpha_k \sim \mathcal{GP}(\mu_k, \mathcal{K}(x_i, x'_i))$$

- \mathcal{GP} depends on parameters $\mu_k, \eta_k, \rho_Z, \rho_N$

Inference procedure

- In total, parameters $\mu_k, \eta_k, \rho_Z, \rho_N, \sigma_k$
- Sample posterior distributions:
joint posterior \propto likelihood \cdot prior
- We employ PyMC¹ for sampling

¹Abril-Pla, O., Andreani, V., Carroll, C., Dong, L., Fannesbeck, C. J., Kochurov, M., Kumar, R., Lao, J.,

Luhmann, C. C., Martin, O. A., et al. (2023). [Pymc: a modern, and comprehensive probabilistic programming framework in python.](#)

PeerJ Computer Science, 9:e1516

Microscopic shell model

- Duflo-Zuker shell model employs two different occupations
 - Natural occupation
 - Particle-hole excitations across HO shell closure
- Intermediate HO excitations of fewer nucleons
- Mixture model of different configurations
 - Particle hole excitations across SO and HO shell closures
- Higher order terms of Hamiltonian
- All fit to binding energy per nucleon for $N, Z \geq 28$

Results - Intermediate HO excitations

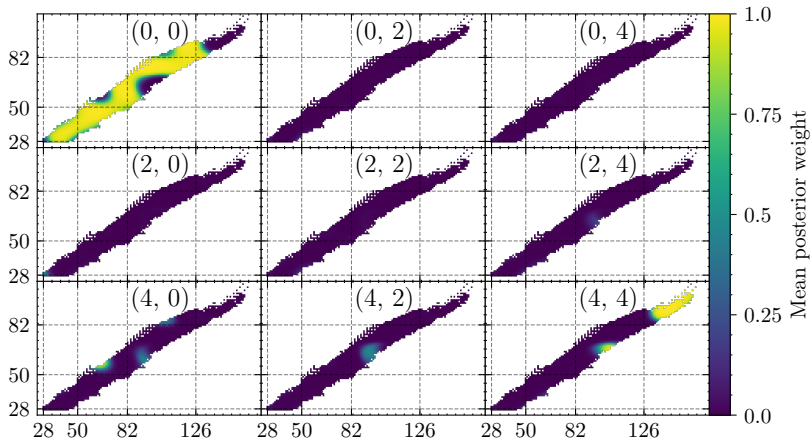


Figure: Posterior weights of mixture with intermediate promotions. Mean RMS: 677 keV, median RMS: 545 keV

Results - SO excitations

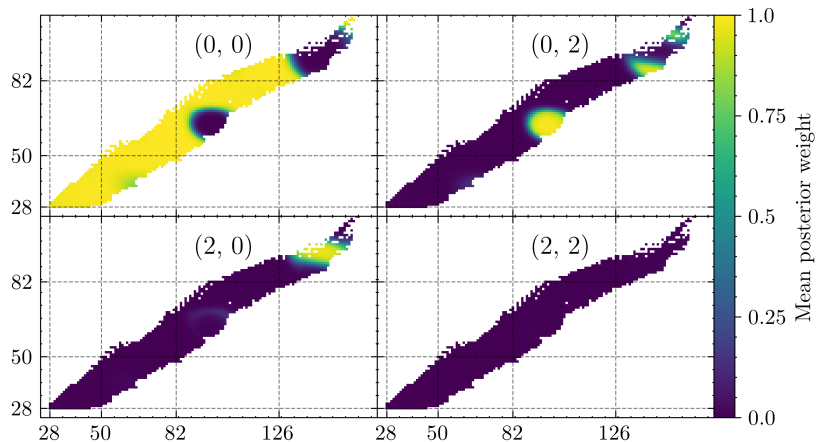


Figure: Posterior weights of mixture with intermediate promotions.

Results - SO promotions

- Mean RMS: 520 keV, median RMS: 505 keV.
- Mean σ_k :
 - (0, 0): 4.4 keV
 - (0, 2): 2.3 keV
 - (2, 0): 2.2 keV
 - (2, 2): 10 keV (corresponds to prior)

Individual uncertainties for each model!

Uncertainty Quantification

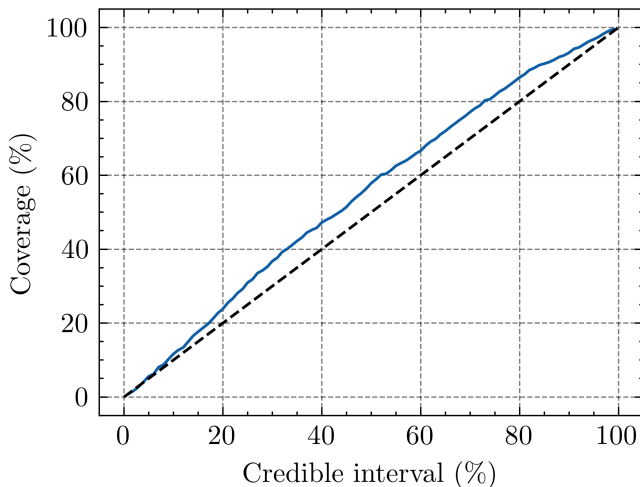


Figure: Credible intervals of posterior distributions vs coverage of experimental data, for the SO promotion model.

Conclusion

- Combination of Bayesian frameworks for local evaluation of models
- Can potentially be used to fit models with location dependent terms
- Demonstration of lower rms by local selection of models
- Uncertainty intervals correlate well with observed model coverage

Thank you!

Questions?

More on model setup

- Prior distributions:

$$\sigma \stackrel{\text{iid}}{\sim} \text{Gamma}(10, 1),$$

$$\mu_k \sim \mathcal{N}(0, 1),$$

$$\eta_k \sim \text{Gamma}(10, 2),$$

$$\rho_Z, \rho_N \stackrel{\text{iid}}{\sim} \text{Gamma}(5, 2)$$

- No-U-Turn Sampler (NUTS) ¹
- Hilbert Space Gaussian Process approximation ² implemented in PyMC.
- Covariance function

$$\mathcal{K}(x_i, x'_j) = \eta_k \exp \left[-\frac{(Z - Z')^2}{2\rho_Z^2} - \frac{(N - N')^2}{2\rho_N^2} \right]$$

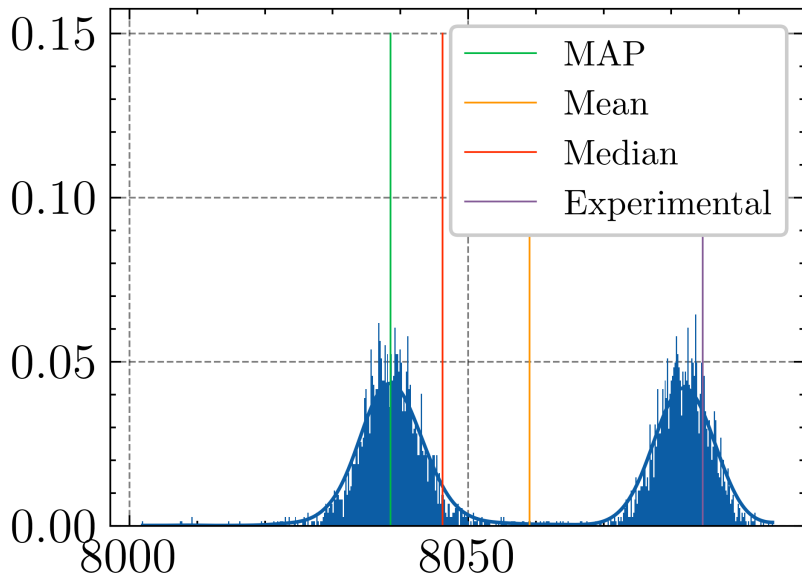
¹Hoffman, M. D., Gelman, A., et al. (2014). [The no-u-turn sampler: adaptively setting path lengths in hamiltonian monte carlo.](#)

J. Mach. Learn. Res., 15(1):1593–1623

²Riutort-Mayol, G., Bürkner, P.-C., Andersen, M. R., Solin, A., and Vehtari, A. (2023). [Practical hilbert space approximate bayesian gaussian processes for probabilistic programming.](#)

Statistics and Computing, 33(1):17

Mean, median and MAP



Inferred existence probabilities

